

Solution of the exercises of the exam of the course

Probability and Stochastic Processes

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Exercise 1 Let $\{U_i\}_{i \geq 1}$ a sequence of i.i.d.r.v.'s on a probability space $(\Omega, \mathcal{F}, \mathbb{P})$ such that $U_1 \stackrel{d}{=} \text{Unif}[0, 1]$. Let \mathcal{F}_0 be the trivial σ algebra and, for any $n \geq 1$, let \mathcal{F}_n be the σ algebra generated by the random vector (U_1, \dots, U_n) . Given $\theta \in (0, 1)$, set $X_0 := p \in (0, 1)$ and $\forall n \geq 1, X_{n+1} := \theta X_n + (1 - \theta) \mathbf{1}_{[0, X_n]}(U_{n+1})$.

1. Prove $\forall n \geq 0, X_n \in [0, 1]$.
2. Prove that $\{X_n\}_{n \geq 0}$ is an L^2 -martingale w.r.t. $\{\mathcal{F}_n\}_{n \geq 0}$.
3. Compute $\mathbb{E}[(X_{n+1} - X_n)^2]$, compute its limit as $n \uparrow \infty$ and deduce the distribution of the limit X_∞ of the terminal value of $\{X_n\}_{n \geq 0}$.

Solution:

1. Since by definition $X_0 \in [0, 1]$, then

$$0 < \theta p \leq X_1 = \theta p + (1 - \theta) \mathbf{1}_{[0, p]}(U_{n+1}) \leq \theta p + (1 - \theta) < 1.$$

Hence, assuming $X_n \in [0, 1]$, we get

$$0 \leq X_{n+1} \leq \theta X_n + (1 - \theta) \leq 1$$

and the statement follows by induction.

2. By the previous statement $\forall n \geq 0, X_n \in L^1$, therefore we are left with the proof of the martingale property. But

$$\begin{aligned} \mathbb{E}[X_{n+1} | \mathcal{F}_n] &= \theta X_n + (1 - \theta) \mathbb{E}[\mathbf{1}_{[0, X_n]}(U_{n+1}) | \mathcal{F}_n] \\ &= \theta X_n + (1 - \theta) X_n = X_n. \end{aligned}$$

Moreover $\forall n \geq 0, \mathbb{E}[X_n^2] \leq 1$ hence Prove that $\{X_n\}_{n \geq 0}$ is an L^2 -martingale.

3.

$$\mathbb{E} \left[(X_{n+1} - X_n)^2 \right] = (1 - \theta)^2 \mathbb{E} \left[(X_n - \mathbf{1}_{[0, X_n]}(U_{n+1}))^2 \right] .$$

But

$$\mathbb{E} \left[(X_n - \mathbf{1}_{[0, X_n]}(U_{n+1}))^2 \right] = \mathbb{E} [X_n^2] - 2\mathbb{E} [X_n \mathbf{1}_{[0, X_n]}(U_{n+1})] + \mathbb{E} [\mathbf{1}_{[0, X_n]}(U_{n+1})]$$

and

$$\begin{aligned} \mathbb{E} [\mathbf{1}_{[0, X_n]}(U_{n+1})] &= \mathbb{E} [\mathbb{E} [\mathbf{1}_{[0, X_n]}(U_{n+1}) | \mathcal{F}_n]] = \mathbb{E} [X_n] , \\ \mathbb{E} [X_n \mathbf{1}_{[0, X_n]}(U_{n+1})] &= \mathbb{E} [\mathbb{E} [X_n \mathbf{1}_{[0, X_n]}(U_{n+1}) | \mathcal{F}_n]] = \\ &= \mathbb{E} [X_n \mathbb{E} [\mathbf{1}_{[0, X_n]}(U_{n+1}) | \mathcal{F}_n]] = \mathbb{E} [X_n^2] , \end{aligned}$$

which implies

$$\mathbb{E} \left[(X_n - \mathbf{1}_{[0, X_n]}(U_{n+1}))^2 \right] = \mathbb{E} [X_n] - \mathbb{E} [X_n^2] = \mathbb{E} [X_n (1 - X_n)]$$

so that

$$\mathbb{E} \left[(X_{n+1} - X_n)^2 \right] = (1 - \theta)^2 \mathbb{E} [X_n (1 - X_n)] .$$

Since $\{X_n\}_{n \geq 0}$ is an L^2 -martingale, it is convergent in L^2 to a \mathcal{F}_∞ -measurable r.v. X_∞ such that $X_n = \mathbb{E}[X_\infty | \mathcal{F}_n]$. Thus

$$\begin{aligned} 0 &\leq (1 - \theta)^2 \lim_{n \uparrow \infty} \mathbb{E} [X_n (1 - X_n)] = \lim_{n \uparrow \infty} \mathbb{E} \left[(X_{n+1} - X_n)^2 \right] \\ &\leq \lim_{n \uparrow \infty} \mathbb{E} \left[((X_{n+1} - X_\infty) + (X_n - X_\infty))^2 \right] \\ &\leq \lim_{n \uparrow \infty} (\|X_{n+1} - X_\infty\|_{L^2} + \|X_n - X_\infty\|_{L^2})^2 = 0 , \end{aligned}$$

which, since X_∞ is non negative, implies that $X_\infty (1 - X_\infty) = 0, \mathbb{P}$ -a.s.. Furthermore, since L^2 -martingales are regular $\{X_n\}_{n \geq 0}$ is convergent to X_∞, \mathbb{P} -a.s. and because $\mathbb{E}[X_n] = \mathbb{E}[X_0] = p > 0$, we get that $X_\infty \stackrel{d}{=} \text{Ber}(p)$.

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Exercise 2 Let $(\Omega, \mathcal{F}, \{\mathcal{F}_t\}_{t \geq 0}, \mathbb{P})$ be the filtered Wiener space. Solve the stochastic process described by the Itô Stochastic Differential Equation

$$\begin{aligned} X(t, X_0) &= X_0 + \int_0^t ds \frac{s}{2} X(s) + \int_0^t 2\sqrt{s} X(s) dB(s) , \\ dX(t) &= \frac{t}{2} X(t) dt + 2\sqrt{t} X(t) dB(t) . \end{aligned} \tag{1}$$

where $(B(t), t \geq 0)$ is the Brownian motion.

1. Compute the variance of the stochastic process $(X(t, 1), t \geq 0)$.

Solution: The equation (1) is a Itô's SDE with multiplicative noise.

1. Setting

$$Y(t) = f(t, X(t)) := \log \frac{X(t)}{X_0} \quad (2)$$

and computing the Itô's differential of $Y(t)$, since $f(t, x) = \log \frac{x}{X_0}$ and

$$\partial_t f(t, x) = 0, \quad (3)$$

$$\partial_x f(t, x) = \frac{X_0}{x}, \quad (4)$$

$$\partial_x^2 f(t, x) = -\frac{X_0}{x^2}, \quad (5)$$

we get

$$dY(t) = \left(\frac{1}{2}t - 2t\right) dt + 2\sqrt{t}dB(t), \quad Y(t) = -\frac{3}{4}t^2 + \int_0^t 2\sqrt{s}dB(s). \quad (6)$$

That is, taking into account that $Y(0, X(0)) = 0$,

$$X(t, X_0) = X_0 e^{-\frac{3}{4}t^2 + \int_0^t 2\sqrt{s}dB(s)}. \quad (7)$$

2. Setting $X_0 = 1$, from (1) we get

$$\mathbb{E}[X(t, 1)] = 1 + \mathbb{E}\left[\int_0^t ds \frac{s}{2} X(s)\right] + \mathbb{E}\left[\int_0^t 2\sqrt{s}X(s)dB(s)\right], \quad (8)$$

that is, $\mu_X(t) := \mathbb{E}[X(t, 1)]$, satisfies the Cauchy's problem

$$\begin{cases} \frac{d}{dt}\mu_X(t) = \frac{t}{2}\mu_X(t) \\ \mu_X(0) = 1 \end{cases}. \quad (9)$$

Hence,

$$\mu_X(t) = e^{\int_0^t ds \frac{s}{2}} = e^{\frac{t^2}{4}}. \quad (10)$$

Moreover, putting $Y(t) := f(t, X(t)) = X^2(t)$ and computing its Itô's differential we have

$$Y(t) = 1 + 5 \int_0^t ds s X^2(s) + \int_0^t 4X^2(s) \sqrt{s}dB(s). \quad (11)$$

Thus, denoting by $q_X(t) := \mathbb{E}[Y(t)] = \mathbb{E}[X^2(t)]$, taking the expectation of both sides of the preceding Itô's equations and taking the derivative w.r.t. t we obtain that q_X is the solution of the Cauchy's problem

$$\begin{cases} \frac{d}{dt}q_X(t) = 5tq_X(t) \\ q_X(0) = 1 \end{cases} \quad (12)$$

i.e.

$$q_X(t) = e^{\frac{5}{2}t^2}. \quad (13)$$

Therefore, the variance of $X(t, 1)$ is

$$\mathbb{E}[X^2(t, 1)] - \mathbb{E}^2[X(t, 1)] = q_X(t) - \mu_X^2(t) = e^{\frac{5}{2}t^2} - e^{\frac{t^2}{2}}. \quad (14)$$

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Exercise 3 Compute the characteristic function of the random vector $(\log Y(1, 1), \log Y(2, 1))$, where $(Y(t, 1), t \geq 0)$ is the stochastic process solution of the homogeneous Itô SDE associated to the equation (1).

Solution: The distribution of the random vector

$$\begin{aligned} Y & : = (\log X(1, 1), \log X(2, 1)) \\ & = \left(-\frac{3}{4} + 2 \int_0^1 \sqrt{t} dB(t), -3 + 2 \int_0^2 \sqrt{t} dB(t) \right) \end{aligned} \quad (15)$$

is Gaussian with parameters

$$\mu = \left(-\frac{3}{4}, -3 \right) \quad (16)$$

and, since

$$a(t) := 4 \int_0^t ds s = 2t^2, \quad (17)$$

covariance matrix

$$C := \begin{pmatrix} 2 & 2 \\ 2 & 8 \end{pmatrix}. \quad (18)$$

Thus, $\forall \lambda = (\lambda_1, \lambda_2) \in \mathbb{R}^2$,

$$\begin{aligned} \varphi_Y(\lambda) & : = \mathbb{E} \left[e^{i\langle \lambda, Y \rangle} \right] = e^{i\langle \mu, \lambda \rangle - \frac{1}{2} \langle C \lambda, \lambda \rangle} \\ & = e^{-i\left(\frac{3}{4}\lambda_1 + 3\lambda_2\right)} \exp \left[-\frac{1}{2} (2\lambda_1^2 + 4\lambda_1\lambda_2 + 8\lambda_2^2) \right] \\ & = e^{-i\left(\frac{3}{4}\lambda_1 + 3\lambda_2\right)} \exp \left[-(\lambda_1^2 + 2\lambda_1\lambda_2 + 4\lambda_2^2) \right]. \end{aligned} \quad (19)$$

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